

One-step validation method for surrogate endpoints in multiple randomized cancer clinical trials with failure-time endpoints

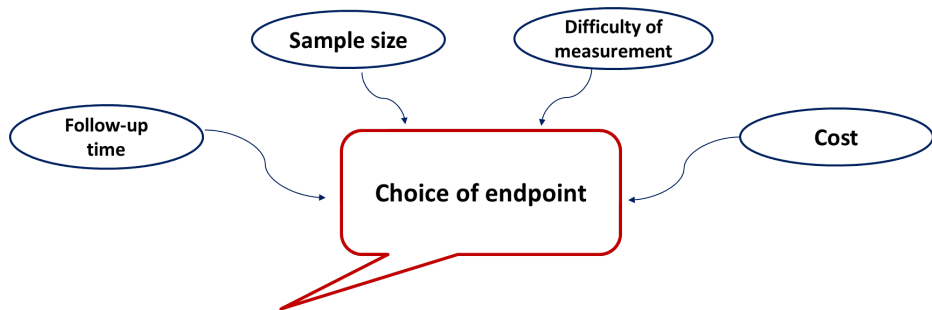
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Setting up clinical trials

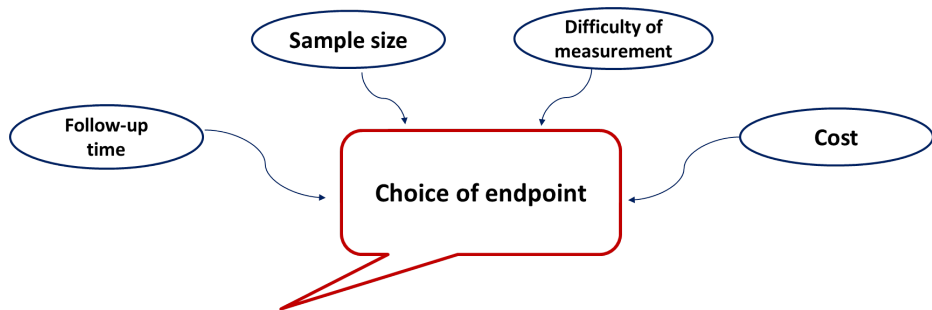


Surrogate endpoint

- Might be measured earlier, more conveniently, or more frequently
- But, need to be validated based on appropriate methods

Burzykowski et al. 2005, Fleming 1996, Ellenberg and Hamilton 1989, Fleming and DeMets 1986

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Surrogate endpoint assessment

Related works: (**S** = surrogate et **T** = True endpoint)

- **S** must capture the full effect of treatment upon the **T**, Prentice (1)
- Validation (1): two-stage model - population and individual level
⇒ Large number of observations, meta-analysis

Current consensus: two failure-time endpoints

- Individual level, association between S and T: Copula model
- Trial-level, prediction: random effect model, adjusted R^2 ($R^2_{trial,adj}$)
- Methods based on a two-step hierarchical analysis strategy

Main problem: $R^2_{trial,adj}$ not always available

Prentice 1989, Freedman et al. 1992, Burzykowski et al. 2001, Renfro et al. 2012, Li et al. 2011

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Main Objectif

To propose a surrogate endpoint one-step validation method, based on a joint frailty model

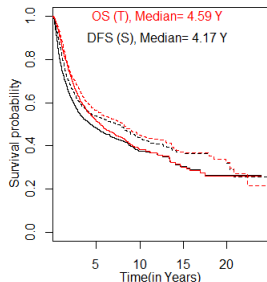
Motivating example

Data source

Individual patient data meta analysis in gastric cancer, from The GASTRIC group: evaluation of an adjuvant therapy, in term of DFS and OS

Data description

- DFS: minimum(relapse, death)
- Included 3288 patients in 14 trials
- Subjects per trial: 6.5 (IQR: 5.5-8.2)%
- Randomized patients: 48%
- 1763 (54%) DFS vs 1705 (52%) OS



Model definition

Assume S_{ij} and T_{ij} , the event-times associated with S and T for subject $j \in \text{trial } i$

$$\begin{cases} \lambda_{S_{ij}}(t|\omega_{ij}, \mathbf{u}_i, \mathbf{v}_{S_i}, \mathbf{Z}_{ij1}) = \lambda_{0S}(t) \exp(\omega_{ij} + \mathbf{u}_i + \mathbf{v}_{S_i} \mathbf{Z}_{ij1} + \beta_S \mathbf{Z}_{ij1}) \\ \lambda_{T_{ij}}(t|\omega_{ij}, \mathbf{u}_i, \mathbf{v}_{T_i}, \mathbf{Z}_{ij1}) = \lambda_{0T}(t) \exp(\zeta \omega_{ij} + \alpha \mathbf{u}_i + \mathbf{v}_{T_i} \mathbf{Z}_{ij1} + \beta_T \mathbf{Z}_{ij1}) \end{cases}$$

where

$$\omega_{ij} \sim N(0, \theta) \quad \mathbf{u}_i \sim N(0, \gamma) \quad \begin{pmatrix} \mathbf{v}_{S_i} \\ \mathbf{v}_{T_i} \end{pmatrix} \sim MVN(\mathbf{0}, \Sigma), \quad \Sigma = \begin{pmatrix} \sigma_S^2 & \sigma_{ST} \\ \sigma_{ST} & \sigma_T^2 \end{pmatrix}$$

- ω_{ij} , heterogeneity: correlation between S and T at the individual-level
- \mathbf{u}_i trial random effect associated with the baseline risk function
- ζ and α , to distinguish heterogeneities between end-points
- \mathbf{v}_{S_i} and \mathbf{v}_{T_i} , correlated random treatment by trial interaction
- \mathbf{v}_{S_i} and \mathbf{v}_{T_i} : assess the prediction of treatment effect on T

Model estimation

Marginal log-likelihood : $l(\Phi)$

$$l(\Phi) = \log \left\{ \prod_{i=1}^G \int_U \left[\prod_{j=1}^{ni} \int_{\omega_{ij}} \lambda_{S_{ij}}^{\delta_{ij}} \cdot S(S_{ij}) \cdot \lambda_{T_{ij}}^{\delta_{ij}^*} \cdot S(T_{ij}) f(\omega_{ij}) d\omega_{ij} \right] f(v_{S_i}, v_{T_i}) f(u_i) dU \right\}$$

where

- $\Phi = (\hat{\sigma}_{v_S}^2, \hat{\sigma}_{v_T}^2, \hat{\sigma}_{v_{ST}}, \hat{\theta}, \hat{\gamma}, \hat{\lambda}_{0T}(\cdot), \hat{\lambda}_{0S}(\cdot), \hat{\beta}_S, \hat{\beta}_T)$
- $U = (u_i, v_{S_i}, v_{T_i})$

Parameters estimation

- 1 $\hat{\lambda}_{0S}(\cdot)$ and $\hat{\lambda}_{0T}(\cdot)$ approximated using splines, smoothing functions
- 2 Maximization of the penalized Log-likelihood using Marquardt algorithm
- 3 Numeric integration with Monte-Carlo and Gaussian Hermite quadrature

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Surrogacy Evaluation and simulations

Individual and Trial-level surrogacy

- New definition of Kendall's $\tau = f(\theta, \zeta, \gamma, \alpha)$
- Trial-level evaluation by $R_{trial}^2 = \frac{\sigma_{ST}^2}{\sigma_S^2 \sigma_T^2}$
- Standard error of R_{trial}^2 by the delta method, bootstrap for Kendall's τ

Simulations

- Data generation using the surrogate joint-model (1)
- Event-times associated to S and T followed a Weibull distribution
- Scale and shape parameters: $\gamma_T = 3.0$, $\rho_T = 0.0025$, $\gamma_S = 1.8$, $\rho_S = 0.0045$
- We assumed S_{ij} censored by T_{ij} with $S_{ij} < T_{ij}$ in case of no censorship

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Integration methods and model assessment

	True	N=600 subjects; G=30 trials			
		Mean	SD	SE	CP
		MC-GH, Cens~40%			
θ	1.000	1.209	0.491	0.505	96
ζ	1.000	0.995	0.374	0.384	90
γ	0.800	0.934	0.390	0.329	93
α	1.000	0.989	0.182	0.204	95
σ_S^2	0.700	0.760	0.414	0.379	90
σ_T^2	0.700	0.730	0.459	0.366	90
σ_{ST}	0.630	0.649	0.352	0.295	90
β_S	-1.250	-1.245	0.273	0.244	91
β_T	-1.250	-1.210	0.241	0.236	91
R_{trial}^2	0.810	0.817	0.186	0.185	75
τ	0.378	0.385	0.037	-	93

$N_{sim} = 500$, $N_{MC} = 300$ samples, $N_{GH} = 20$ points

1 Integration assessment

- Good parameters' estimation
- MC, PGH, MC-GH, MC-PGH
- Comparable results found

2 Variation of data parameters

- $N=1000$, $G=10$, Cens=70%
- Robust. evaluation criteria
- Improved CP of R^2 when $N \uparrow$

Comparison with two-steps approach

Parameters	True	600 subjects; 30 trials			NC:n(%)	600 subjects; 10 trials			NC:n(%)
		Mean	Bias	MSE		Mean	Bias	MSE	
High individual-level and trial-level associations									
Clayton [†]					266(53)				10(2)
Kendall's τ	0.614	0.565	0.049	0.003		0.543	0.071	0.006	
$R^2_{trial,adj}$	0.810	0.742	0.068	0.082		0.833	-0.023	0.061	
Plackett [†]					274(55)				14(3)
Kendall's τ	0.614	0.559	0.055	0.004		0.538	0.076	0.006	
$R^2_{trial,adj}$	0.810	0.811	-0.001	0.069		0.844	-0.034	0.062	
Joint Surrogate [‡]					1(0)				4(1)
Kendall's τ	0.614	0.613	0.001	0.002		0.614	<10-3	0.001	
R^2_{trial}	0.810	0.817	-0.007	0.057		0.816	-0.006	0.061	
Weak individual-level and trial-level associations									
Clayton [†]					236(47)				8(2)
Kendall's τ	0.378	0.277	0.101	0.011		0.257	0.121	0.016	
$R^2_{trial,adj}$	0.360	0.446	-0.086	0.105		0.525	-0.165	0.139	
Plackett [†]					238(48)				10(2)
Kendall's τ	0.378	0.278	0.100	0.011		0.271	0.107	0.012	
$R^2_{trial,adj}$	0.360	0.511	-0.151	0.124		0.512	-0.152	0.139	
Joint Surrogate [‡]					0(0)				0(0)
Kendall's τ	0.378	0.382	-0.004	0.001		0.375	0.003	0.004	
R^2_{trial}	0.360	0.442	-0.082	0.08		0.478	-0.118	0.106	

[†] Two-step approaches, [‡] One-step approaches, NC = did Not converged, N_sim = 500 simulated dataset

- Convergence issues with the two-step approaches
- underestimation of Kendall's τ

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- Better estimation of the R^2_{trial} by the proposed approach
- High bias on R^2_{trial} in case of weak trial level association

Evaluation of DFS as surrogate endpoint for OS

Questions

- 1 Are DFS and OS correlated at the individual-level? Kendall's τ
- 2 Does the treatment effect on DFS reliably predict the treat. effect on OS?, R_{trial}^2

DFS Evaluation

Parameters	R_{trial}^2	Kendall's τ
One step approach		
Proposed joint surrogate	0.99 (0.84-1.14)	0.68 (0.65-0.70)
Poisson	1.00 (0.08-1.00)	0.74 (0.73-0.76)
Two steps approach using R		
Clayton adj	0.97 (0.46-1.00)	0.81 (0.80-0.91)
Plackett adj	1.00 (0.69-1.00)	0.82 (0.81-0.83)
Hougaard adj	0.94 (0.08-1.00)	0.18 (0.17-0.19)

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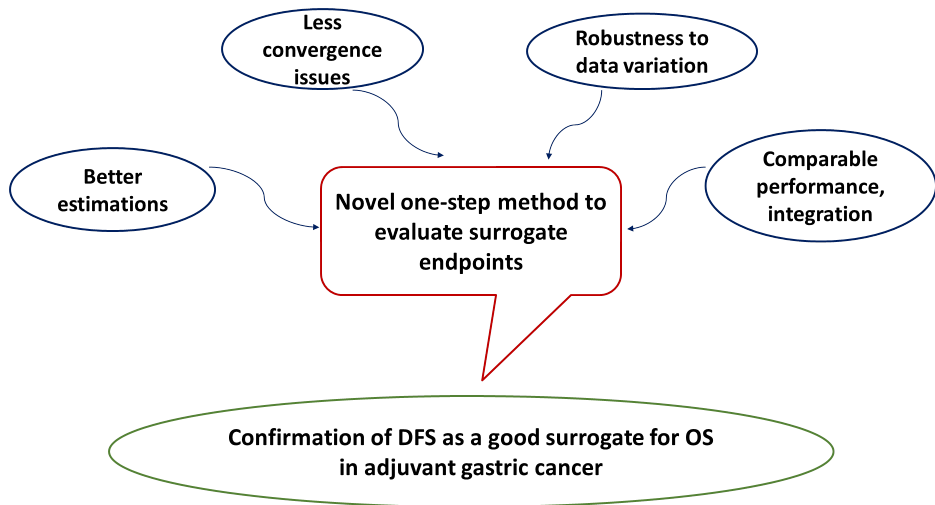
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Conclusion



Funding :



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Reconnue d'utilité publique



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Thanks for your kind attention

The semi-parametric penalized likelihood approach

Parameters estimation

- The baseline hazard function were approximated using cubic M-splines
- $\sigma_{v_S}^2, \sigma_{v_T}^2, \sigma_{v_{ST}}, \theta, \gamma, \beta^S, \beta^T, \lambda_{0S}(t)$ and $\lambda_{0T}(t)$ estimated based on a semi parametric penalized likelihood approach.
- The penalized log-likelihood were defined by:

$$pl(\Phi) = l(\Phi) - k_1 \int_0^\infty \lambda_{0S}''^2(t) dt - k_2 \int_0^\infty \lambda_{0T}''^2(t) dt$$

- The smoothing parameters k_1 and k_2 were obtained by cross validation
- Estimated parameters were obtained by the robust Marquardt algorithm
- Transformation applied to spline coefficient, variance parameters, ≥ 0
- Standard errors of transformed parameters based on Delta-method

Individual-level surrogacy

- New definition of Kendall's τ to evaluate the individual level surrogacy.
- $\tau \in [-1,1]$, assumes a zero value when S_{ij} and T_{ij} are independent.
- One can show that

$$\tau = 2 \int_{u_i} \int_{\omega_{ij}} \int_{u_{i'}} \int_{\omega_{i'j'}} \frac{\exp(\omega_{ij} + u_i + \zeta\omega_{ij} + \alpha u_i) + \exp(\omega_{i'j'} + u_{i'} + \zeta\omega_{i'j'} + \alpha u_{i'})}{(\exp(\omega_{i'j'} + u_{i'}) + \exp(\omega_{ij} + u_i))(\exp(\zeta\omega_{i'j'} + \alpha u_{i'}) + \exp(\zeta\omega_{ij} + \alpha u_i))} \frac{1}{\sqrt{2\pi\theta}} \exp\left[-\frac{1}{2} \frac{\omega_{i'j'}^2}{\theta}\right] \frac{1}{\sqrt{2\pi\gamma}} \exp\left[-\frac{1}{2} \frac{u_{i'}^2}{\gamma}\right] d\omega_{i'j'} du_{i'} \frac{1}{\sqrt{2\pi\theta}} \exp\left[-\frac{1}{2} \frac{\omega_{ij}^2}{\theta}\right] \frac{1}{\sqrt{2\pi\gamma}} \exp\left[-\frac{1}{2} \frac{u_i^2}{\gamma}\right] d\omega_{ij} du_i - 1$$

α and ζ set to 1, n_i and randomized subject per trial variables

Parameters	True value	Mean	Empirical SE	Mean SE	CP(%)
		N=1000 subjects G=14 trials C=48%			
θ	3.500	3.407	0.638	0.343	88
γ	2.500	2.733	1.994	0.651	59
σ_S^2	0.700	0.663	0.451	0.338	79
σ_T^2	0.700	0.669	0.451	0.350	77
σ_{ST}	0.630	0.593	0.417	0.309	77
β_S	-1.250	-1.211	0.328	0.250	88
β_T	-1.250	-1.208	0.333	0.259	85
R_{trial}^2	0.810	0.816	0.217	0.189	69
Kendall's τ	0.584	0.583	0.055	-	69
NC : n(%)	-	6(1)	-	-	-

Initial values quite far from true simulation values

Parameters	True value	Mean	Empirical SE	Mean SE	CP(%)
		N=1000 subjects G=30 trials C=38%			
θ	3.500	3.474	0.808	0.609	84
ζ	1.000	1.110	0.456	0.187	84
γ	2.500	2.627	0.933	0.640	84
α	1.000	1.058	0.230	0.139	86
σ_S^2	0.700	0.692	0.420	0.354	84
σ_T^2	0.700	0.716	0.477	0.390	89
σ_{ST}	0.630	0.618	0.373	0.321	86
β_S	-1.250	-1.177	0.291	0.219	87
β_T	-1.250	-1.190	0.314	0.239	91
R_{trial}^2	0.810	0.811	0.196	0.204	80
Kendall's τ	0.597	0.595	0.029	-	84
NC : n(%)	-	0(0)	-	-	-

Parameters	True	N=600 subjects; G=30 trials				N=600 subjects; G=30 trials			
		Mean	SD	SE	CP	Mean	SD	SE	CP
		Full MC				Full PGH			
θ	1.000	1.122	0.305	0.242	97	1.025	0.418	0.370	89
ζ	1.000	0.990	0.367	0.379	91	1.385	1.490	0.373	87
γ	0.800	0.957	0.493	0.343	92	0.762	0.255	0.273	91
α	1.000	0.982	0.185	0.202	95	1.082	0.310	0.192	92
σ_S^2	0.700	0.787	0.540	0.397	90	0.638	0.341	0.319	83
σ_T^2	0.700	0.723	0.450	0.361	89	0.726	0.495	0.360	84
σ_{ST}	0.630	0.652	0.355	0.296	89	0.581	0.293	0.287	87
β_S	-1.250	-1.259	0.312	0.250	91	-1.253	0.249	0.222	92
β_T	-1.250	-1.205	0.236	0.234	92	-1.305	0.345	0.236	91
R^2_{trial}	0.810	0.817	0.186	0.183	76	0.829	0.195	0.139	55
Kendall's τ	0.378	0.382	0.039	-	90	0.369	0.041	-	88
		MC-GH				MC-PGH			
θ	1.000	1.209	0.491	0.505	96	1.228	0.552	0.529	95
ζ	1.000	0.995	0.374	0.384	90	1.020	0.644	0.381	90
γ	0.800	0.934	0.390	0.329	93	0.939	0.402	0.336	93
α	1.000	0.989	0.182	0.204	95	0.985	0.188	0.204	95
σ_S^2	0.700	0.760	0.414	0.379	90	0.766	0.428	0.385	90
σ_T^2	0.700	0.730	0.459	0.366	90	0.739	0.499	0.367	89
σ_{ST}	0.630	0.649	0.352	0.295	90	0.653	0.355	0.297	90
β_S	-1.250	-1.245	0.273	0.244	91	-1.247	0.280	0.247	91
β_T	-1.250	-1.210	0.241	0.236	91	-1.208	0.241	0.236	91
R^2_{trial}	0.810	0.817	0.186	0.185	75	0.818	0.186	0.184	76
Kendall's τ	0.378	0.385	0.037	-	93	0.384	0.039	-	92

N_sim = 350 simulated dataset, N_MC = 300 samples, N_FullPGH = 12 points and 20 points elsewhere

Parameters	True	Mean	SD	SE	CP	Mean	SD	SE	CP		
		Reference(N=600, G=30, C= \sim 40%)					G=10 trials				
θ	3.500	3.397	0.528	0.526	94	3.372	0.456	0.534	95		
ζ	1.500	1.583	0.324	0.319	85	1.594	0.318	0.337	88		
γ	2.500	2.666	0.971	0.753	88	2.617	1.644	0.834	68		
α	1.000	1.041	0.192	0.187	87	1.048	0.197	0.203	90		
σ_S^2	0.700	0.731	0.518	0.481	88	0.714	0.601	0.442	81		
σ_T^2	0.700	0.850	0.753	0.704	91	0.820	0.791	0.648	84		
σ_{ST}	0.630	0.690	0.550	0.515	90	0.672	0.613	0.470	84		
β_S	-1.250	-1.183	0.274	0.255	92	-1.197	0.351	0.314	91		
β_T	-1.250	-1.229	0.386	0.360	94	-1.227	0.450	0.419	92		
R_{trial}^2	0.810	0.826	0.232	0.356	71	0.814	0.254	0.307	64		
Kendall's τ	0.614	0.614	0.025	-	91	0.612	0.032	-	82		
NC* : n(%)	-	1(0)	-	-	-	4(1)	-	-	-		
		N=1000 subjects					Censoring = 70%				
θ	3.500	3.547	0.629	0.525	93	3.777	0.850	1.124	97		
ζ	1.500	1.601	1.026	0.271	85	1.577	1.101	0.651	91		
γ	2.500	2.591	0.819	0.629	89	2.832	1.117	0.995	92		
α	1.000	1.016	0.210	0.154	86	0.989	0.211	0.353	93		
σ_S^2	0.700	0.708	0.433	0.367	87	0.817	0.696	0.628	88		
σ_T^2	0.700	0.733	0.572	0.478	87	0.952	1.110	1.023	86		
σ_{ST}	0.630	0.627	0.401	0.369	88	0.738	0.679	0.651	88		
β_S	-1.250	-1.198	0.244	0.218	91	-1.230	0.322	0.323	94		
β_T	-1.250	-1.199	0.294	0.292	93	-1.237	0.417	0.502	95		
R_{trial}^2	0.81	0.812	0.206	0.259	78	0.806	0.264	0.858	71		
Kendall's τ	0.614	0.613	0.041	-	91	0.617	0.047	-	93		
NC* : n(%)	-	1(0)	-	-	-	52(10)	-	-	-		

* NC =did Not converged, N_sim = 500 simulated dataset, N_{MC} = 300 samples, N_{GH} = 20 or 32 points