

# New Model-Based Bioequivalence Statistical Approaches for Pharmacokinetic Studies with Sparse Sampling

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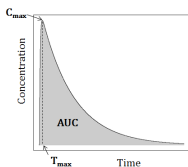
<sup>3</sup> Food and Drug Administration, USA

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# Bioequivalence (BE) studies

- **BE:** The difference in the pharmacokinetic (PK) of two formulations of a given drug does not exceed a predefined threshold (usually  $\delta = \log(1.25)$ )
- **Parallel studies** preferable for drugs with **long half-life:**  $N/2$  subjects receive reference treatment (R),  $N/2$  subjects receive test treatment (T)
- **Non-Compartmental Approach (NCA) based BE:**

- ▶ Individual NCA estimates of  $AUC_i$  and  $C_{max_i}$
- ▶ Let  $\beta_{AUC} = \text{mean}(\log(AUC_{iT})) - \text{mean}(\log(AUC_{iR}))$   
 $\beta_{C_{max}} = \text{mean}(\log(C_{max}_{iT})) - \text{mean}(\log(C_{max}_{iR}))$



pros	cons
<ul style="list-style-type: none"> <li>★ Reproducible</li> <li>★ Few assumptions</li> </ul>	<ul style="list-style-type: none"> <li>★ <b>Require more than 10 samples per subject</b></li> <li>★ Not appropriate for complex models</li> </ul>

- **Two one-sided tests (TOST)**<sup>1</sup> at level  $\alpha = 5\%$  on  $\beta_{AUC}$  and  $\beta_{C_{max}}$ :  
 Reject of  $H_0$  (= BE significant) if
  - ▶  $(\beta - \delta) / SE(\beta) \leq -z_{1-\alpha}$  and  $(\beta + \delta) / SE(\beta) \geq z_{1-\alpha}$ , with  $z_{1-\alpha}$ :  $(1 - \alpha)\%$  quantile of the normal distribution **or**
  - ▶  $CI(\beta)_{1-2\alpha\%} \in [-\delta; \delta]$  (CI: Confidence Interval)

<sup>1</sup>Schuirmann, *J Pharmacokinet Pharmacodyn*, 1987.

# Model based (MB) BE

**Non linear mixed effect model (NLMEM):**

$y_{ij}$ : concentration for subject  $i$  at sampling time  $j$

$$y_{ij} = f(t_{ij}, \phi_i) + g(t_{ij}, \phi_i) \epsilon_{ij}$$

$$\log(\phi_{il}) = \log(\lambda_l) + \beta_l Tr_i + \eta_{il} \text{ where}$$

- ▶  $\beta_l$ : Test treatment effect on the log of a PK parameter  $l = (1, \dots, p)$
- ▶  $\lambda_l$ :  $l^{th}$  element of the vector of fixed effects
- ▶  $Tr_i$ : vector of indicator for treatment group
- ▶  $\eta_{il} \sim \mathcal{N}(0, \omega_l)$ : between subject random effect for parameter  $l$
- ▶  $\epsilon_{ij} \sim \mathcal{N}(0, 1)$ : residual error
- ▶ combined error model  $g(t_{ij}, \phi_i) = a + b \times f(t_{ij}, \phi_i)$

Vector of population parameters:  $\theta = (\lambda, \beta, \omega, a, b)$

- Estimation of  $\theta$  using SAEM algorithm<sup>2</sup>
- Estimation of  $SE(\theta)$  from observed Fisher information matrix  $\Sigma$

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<sup>2</sup>Kuhn et Lavielle. *Comput Stat Data Anal*, 2005.

# MBBE

- AUC and Cmax: secondary parameters of the NLMEM
  - ▶  $\beta_{AUC}$  and  $\beta_{Cmax}$  functions of vectors  $\lambda$  and  $\beta$
  - ▶  $SE(\beta_{AUC})$  and  $SE(\beta_{Cmax})$  obtained using the delta method<sup>3</sup>
- TOST at level  $\alpha = 5\%$  or 90%CI on  $\beta_{AUC}$  and  $\beta_{Cmax}$

pros	cons
★ Require few samples per subject	★ SE under estimated on sparse designs ⇒ Type I error inflation <sup>4</sup>

- Objectives: Development, evaluation and comparison of new model-based (MB) statistical approaches for sparse design BE studies
  - ▶ Parametric random effect and residual bootstrap
  - ▶ Full distribution estimation using Stan<sup>5</sup>

<sup>3</sup>Oelhert *The American Statistician*, 1992.

<sup>4</sup>Dubois et al. *Stat in Med*, 2011.

<sup>5</sup>Stan development team, Rstan, 2012.

# MBBE parametric random effect and residual bootstrap

## Principle<sup>6</sup>

- 1 Estimation of  $\theta$  and  $\Sigma$  with saemix
- 2 Drawing of  $b = 1, \dots, B$  ( $B=250$ ) matrices of random effects of size  $N \times p$  from  $\mathcal{N}(0, \hat{\Omega})$
- 3 Drawing of vector of residual errors of size  $\sum_{i=1}^N n_i$  from  $\mathcal{N}(0, 1)$
- 4 Simulation of the  $B$  vectors of responses
- 5 Fit the  $B$  new data sets with saemix to get the  $B$  estimates  $\theta_b$  and  $\beta_b$
- 6 Derive 90% CIs on  $\beta_{Cmax}$  and  $\beta_{AUC}$  from the 5th and 95th percentiles of the serie  $\hat{\beta}_b$

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<sup>6</sup>Thai et al, *J Pharmacokinet Pharmacodyn*, 2014.

# MBBE full distribution using Stan

## Principle<sup>7</sup>

- 1 Estimation of  $\theta$  and  $\eta_i$  with saemix
- 2 Full distribution in Stan<sup>5</sup>
  - ▶ Initialize HMC chain at estimates from step 1
  - ▶ Draw 1 chain of 1000 (including 100 burning) samples in *a posteriori* distributions of  $\lambda, \beta, \omega, a, b$
  - ▶ out of the  $B=900$  samples derive 90%CI on  $\beta_{Cmax}$  and  $\beta_{AUC}$

Stan model:

- Default distributions on fixed effects  $\lambda, \beta$
- Non-informative priors on  $\omega, a, b$ 
  - $\text{omega\_vec} \sim \text{cauchy}(0, 2.5);$
  - $a \sim \text{cauchy}(0, 2.5);$
  - $b \sim \text{cauchy}(0, 2.5);$

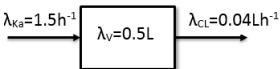
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<sup>5</sup>Stan development team, Rstan, 2012.

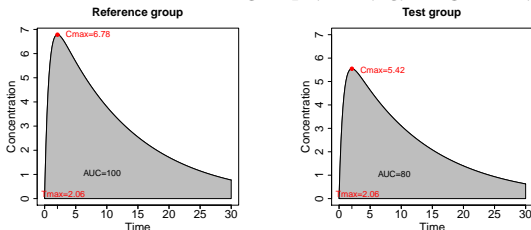
<sup>7</sup>Ueckert et al. ACOP, 2015.

# PK model

- PK model of concentrations of the anti-asthmatic drug theophylline<sup>2</sup>
- Parallel design:  $N_R = N_T = 20$
- One-compartment model with first-order absorption and first-order elimination. Dose=4mg



- Covariate effects in test treatment group:  $\beta_V = \beta_{CL} = \log(1.25)$ ,  $\beta_{K_a} = 0$



- Residual variability:  $a = 0.1$  mg/L,  $b = 10\%$
- Random effects for BSV only:

$\omega_{k_a}$ (%)	$\omega_{V/F}$ (%)	$\omega_{CL/F}$ (%)
22	11	22

# Simulation scenarios

- 500 simulated data sets for each of the 4 scenarios

Design	N	Sampling times (hours)	Hypothesis	$\beta_{CL} = \beta_V$
Rich	40	n=10, $t = (0.25, 0.5, 1, 2, 3.5, 5, 7, 9, 12, 24)$	$H_0$	$\log(1.25)$
			$H_1$	$\log(1)$
Sparse	40	n=3, $t = (0.25, 3.35, 24)$	$H_0$	$\log(1.25)$
			$H_1$	$\log(1)$

Under  $H_0$ :  $\beta_{AUC} = \beta_{Cmax} = \log(0.8)$ ,

Under  $H_1$ :  $\beta_{AUC} = \beta_{Cmax} = \log(1) = 0$

- Evaluation:
  - ▶ Type I error ( $IP_{95\%}(0.05) = [0.033; 0.073]$ )
  - ▶ Power
  - ▶ Computing times



# Type I error and power estimates

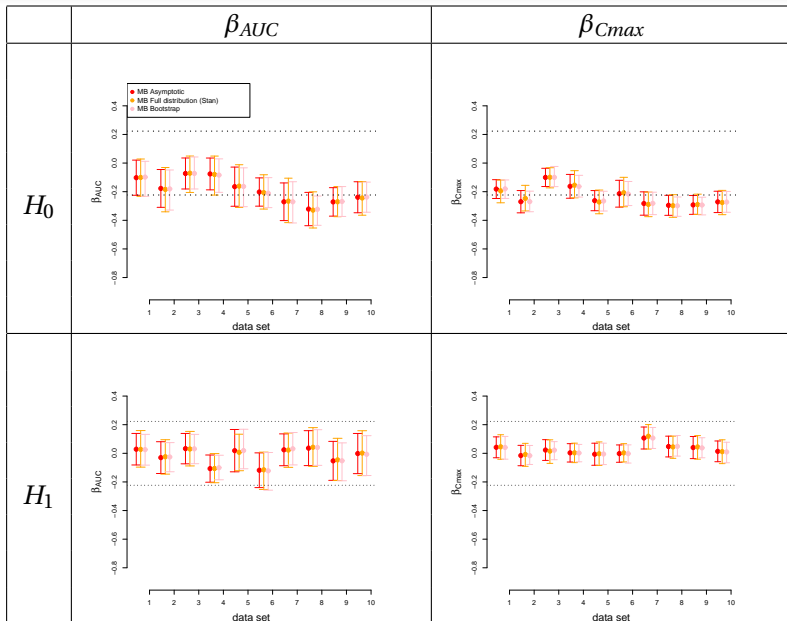
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Mean computing time for 1 data set	15 sec	15 sec	43 min	5 min

	Rich design			
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Type I error $\beta_{AUC}$	0.058	0.056	0.062	0.040
Type I error $\beta_{Cmax}$	0.062	0.058	0.064	0.044
Power $\beta_{AUC}$	0.814	0.830	0.832	0.762
Power $\beta_{Cmax}$	0.998	1.000	1.000	0.962

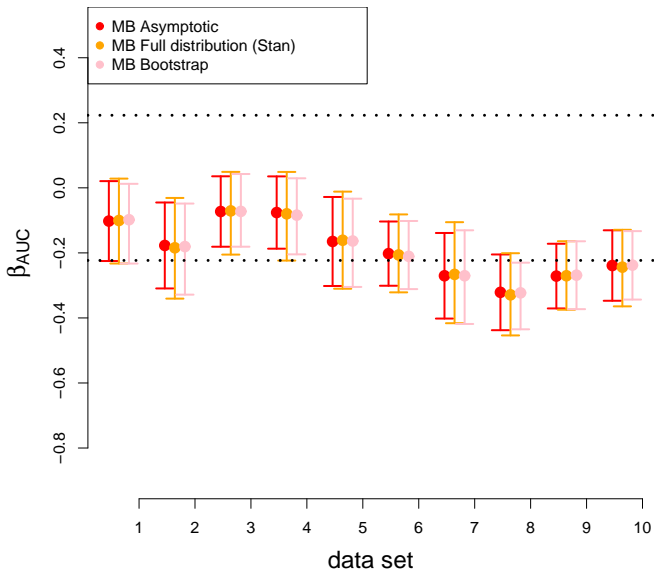
	Sparse design			
	NCA	MB Asymptotic SE	MB Bootstrap CI	MB Full distribution CI
Type I error $\beta_{AUC}$	-	<b>0.076</b>	<b>0.074</b>	0.060
Type I error $\beta_{Cmax}$	-	0.066	0.068	0.050
Power $\beta_{AUC}$	-	0.910	0.916	0.868
Power $\beta_{Cmax}$	-	1.000	1.000	0.992

$IP_{95\%}(0.05) = [0.033; 0.073]$  with 500 simulated data sets

# 90% CI( $\beta$ ) on sparse design



# 90% CI( $\beta$ ) on sparse design, under H0

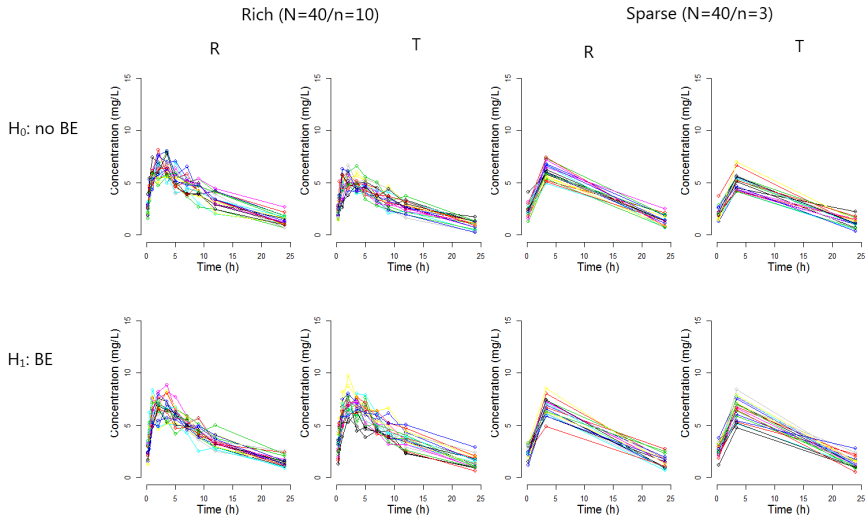


# Conclusion

- Implementation of new methods for MBBE using parametric bootstrap and full distribution sampled in Stan
- Correction of MB TOST with asymptotic SE on sparse design using full distribution sampled in Stan
  - ▶ faster than bootstrap
- Perspective: implementation of the methods for 2-period 2-sequence crossover BE studies (accounting for within-subjects variability)

**Thank you**

# Simulation scenarios



# 90% CI( $\beta$ ) on rich design

